

AN INTRODUCTION TO FINANCIAL MATHEMATICS

YURI KIFER

1. LECTURE 2: STOPPING TIMES AND MARTINGALE INEQUALITIES

1.1. Optional sampling theorem.

1.1. Definition. Given a filtration $\{\mathcal{F}_t, t \geq 0\}$ (or $\{\mathcal{F}_n, n \in \mathbb{Z}_+\}$) a nonnegative random variable τ is called a stopping time if $\{\tau \leq t\} \in \mathcal{F}_t$ for any t . In discrete time case this is equivalent to say that $\{\tau = n\} \in \mathcal{F}_n$ for any n .

If τ_1 and τ_2 are stopping times then so are $\tau_1 \vee \tau_2 = \max(\tau_1, \tau_2)$ and $\tau_1 \wedge \tau_2 = \min(\tau_1, \tau_2)$.

Example. $\tau_B = \min\{n \geq 0 : X_n \in B\}$ is a stopping time for any Borel set B provided X_n is \mathcal{F}_n -measurable for each n .

1.2. Lemma. Let $\{M_n\}_{n \geq 0}$ be a martingale (submartingale) with respect to a filtration $\{\mathcal{F}_n\}$ and τ is a stopping time. Then $M_n^\tau = M_{n \wedge \tau}$ is a martingale (submartingale).

Proof. We write

$$M_{n \wedge \tau} = \sum_{m=0}^{n-1} M_m \mathbb{I}_{\tau=m} + M_n \mathbb{I}_{\tau \geq n}$$

where $\mathbb{I}_\Gamma = 1$ if an event Γ occurs and $= 0$ for otherwise. Hence,

$$M_{(n+1) \wedge \tau} - M_{n \wedge \tau} = \mathbb{I}_{\tau > n} (M_{n+1} - M_n),$$

and so $E(M_{(n+1) \wedge \tau} - M_{n \wedge \tau} | \mathcal{F}_n) = 0$ in the martingale case or $E(M_{(n+1) \wedge \tau} - M_{n \wedge \tau} | \mathcal{F}_n) \geq 0$ in the submartingale case. \square

1.3. Definition. If τ is a stopping time and $\{\mathcal{F}_t\}_{t \geq 0}$ is a filtration. Then $\mathcal{F}_\tau = \{A \in \mathcal{F} : A \cap \{\tau \leq t\} \in \mathcal{F}_t, \forall t\}$.

Clearly, τ is \mathcal{F}_τ -measurable.

1.4. Theorem. (Optional sampling theorem) Let $X = (X_n, \mathcal{F}_n)_{n \geq 0}$ (where $X_n, n \geq 0$ is a sequence of random variables and $\{\mathcal{F}_n\}_{n \geq 0}$ is a filtration) be a martingale (submartingale). Let $\tau_1 \leq \tau_2$ be stopping times such that $E|X_{\tau_i}| < \infty, i = 1, 2$ and

$$\liminf_{n \rightarrow \infty} \int_{\{\tau_2 > n\}} |X_n| dP = 0.$$

Then $E(X_{\tau_2} | \mathcal{F}_{\tau_1}) = X_{\tau_1}$ in the martingale case and $E(X_{\tau_2} | \mathcal{F}_{\tau_1}) \geq X_{\tau_1}$ in the submartingale case.

Date: September 7, 2016.

Proof. It suffices to prove that for any $A \in \mathcal{F}_{\tau_1}$,

$$\int_A X_{\tau_2} dP = \int_A X_{\tau_1} dP$$

in the martingale case and change $=$ to \geq in the submartingale case. Then it suffices to prove that for any $n \in \mathbb{Z}_+$,

$$\int_{A \cap \{\tau_1=n\}} X_{\tau_2} dP = \int_{A \cap \{\tau_1=n\}} X_n dP$$

in the martingale case and change $=$ to \geq in the submartingale case. Now,

$$\begin{aligned} \int_{A \cap \{\tau_1=n\}} X_n dP &= \int_{A \cap \{\tau_1=n\} \cap \{\tau_2 \geq n\}} X_n dP = \int_{A \cap \{\tau_1=n\} \cap \{\tau_2=n\}} X_n dP \\ &+ \int_{A \cap \{\tau_1=n\} \cap \{\tau_2 > n\}} X_n dP = (\text{or } \leq \text{ in the submartingale case}) \int_{A \cap \{\tau_1=n\} \cap \{\tau_2=n\}} X_{\tau_2} dP \\ &\quad + \int_{A \cap \{\tau_1=n\} \cap \{\tau_2 > n\}} E(X_{n+1} | \mathcal{F}_n) dP = \int_{A \cap \{\tau_1=n\} \cap \{\tau_2=n\}} X_{\tau_2} dP \\ &\quad + \int_{A \cap \{\tau_1=n\} \cap \{\tau_2 \geq n+1\}} X_{n+1} dP = \dots = (\text{or } \leq \dots \leq \text{ for submartingales}) \\ &\quad \int_{A \cap \{\tau_1=n\} \cap \{n \leq \tau_2 \leq m\}} X_{\tau_2} dP + \int_{A \cap \{\tau_1=n\} \cap \{\tau_2 > m\}} X_m dP. \end{aligned}$$

Letting $m \rightarrow \infty$ and using the conditions we derive the result. \square

1.5. Definition. A sequence $\{X_n\}$ of random variables is uniformly integrable if

$$\lim_{a \rightarrow \infty} \sup_n \int_{|X_n| > a} |X_n| dP = 0.$$

1.6. Corollary. If $\{X_n\}$ is a uniformly integrable martingale or submartingale then the conditions of the lemma are satisfied, and so its result holds true, as well. In particular, this is true when $|X_n| \leq C$ for all n and some constant C .

1.7. Corollary. If $0 \leq \tau_i \leq N$, $i = 1, 2$ and $X_n, n \geq 0$ is a martingale then

$$EX_0 = EX_{\tau_1} = EX_{\tau_2} = EX_N.$$

1.2. Martingale inequalities.

1.8. Theorem. (see [2]) Let $X_t, t \geq 0$ be a submartingale and $0 \leq t_1 < t_2 < \dots < t_n$.

(i) Set $X = \max_i X_{t_i}$ and $a > 0$. Then

$$aP\{X \geq a\} \leq EX_{t_n} \mathbb{1}_{\{X \geq a\}} \leq EX_{t_n}^+ \leq E|X_{t_n}|.$$

In particular, if X_t is a martingale and $p \geq 1$ then

$$P\{\max_i |X_{t_i}| \geq a\} \leq \frac{E|X_{t_n}|^p}{a^p}.$$

(ii) Suppose that $X_t \geq 0$ is a submartingale and $p > 1$. Then

$$E(\max_i X_{t_i}^p) \leq \left(\frac{p}{p-1}\right)^p EX_{t_n}^p.$$

In particular, if X_t is a martingale (not assuming nonnegativity) then

$$E(\max_i |X_{t_i}|^p) \leq \left(\frac{p}{p-1}\right)^p E|X_{t_n}|^p.$$

Proof. (i) Define the stopping time

$$\tau = \begin{cases} \min\{t_i : X_{t_i} \geq a\} & \text{if an } i \text{ exists that the event in brackets occurs} \\ t_n & \text{otherwise} \end{cases}$$

Then

$$EX_{t_n} \geq EX_\tau = EX_\tau \mathbb{I}_{X \geq a} + EX_{t_n} \mathbb{I}_{X < a} \geq aP\{X \geq a\} + EX_{t_n} \mathbb{I}_{X < a},$$

and so

$$EX_{t_n} \mathbb{I}_{X \geq a} \geq aP\{X \geq a\}$$

implying the assertion (i).

(ii) We have

$$\begin{aligned} EX^p &= \int_\Omega dP \int_0^X p\lambda^{p-1} d\lambda = \int_\Omega dP \int_0^\infty \mathbb{I}_{\{X \geq \lambda\}} p\lambda^{p-1} d\lambda \\ &= p \int_0^\infty \lambda^{p-1} P\{X \geq \lambda\} d\lambda \leq p \int_0^\infty \lambda^{p-2} (EX_{t_n} \mathbb{I}_{\{X \geq \lambda\}}) d\lambda \\ &= p \int_0^\infty \int_\Omega \lambda^{p-2} \mathbb{I}_{\{X \geq \lambda\}} X_{t_n} dP d\lambda = \frac{p}{p-1} \int_\Omega X^{p-1} X_{t_n} dP \end{aligned}$$

where we used (i) and integrated

$$\int_0^\infty \lambda^{p-2} \mathbb{I}_{\{X \geq \lambda\}} d\lambda = \int_0^X \lambda^{p-2} d\lambda = \frac{X^{p-1}}{p-1}.$$

By the Hölder inequality $(E|YZ|) \leq (E|Y|^p)^{1/p} (E|Z|^q)^{1/q}$, $p, q > 0, 1/p + 1/q = 1$,

$$EX^{p-1} X_{t_n} \leq (EX^p)^{\frac{p-1}{p}} (EX_{t_n}^p)^{1/p}.$$

It follows that

$$EX^p \leq \frac{p}{p-1} (EX^p)^{\frac{p-1}{p}} (EX_{t_n}^p)^{1/p},$$

and so

$$\left(\frac{p}{p-1}\right)^p EX_{t_n}^p \geq EX^p$$

completing the proof. \square

Remark Another important result about submartingales is the Doob upcrossing inequality which yields the (sub)martingale convergence theorem (see [1]) but we will not need it in this course.

REFERENCES

- [1] P. Billingsley, *Probability and Measure*, 2nd ed., J. Wiley, New York, 1986.
- [2] A.N. Shiryaev, *Probability*, Springer, New York, 1984.
- [3] D. Williams, *Probability with Martingales*, Cambridge Univ. Press, Cambridge, 1991.